

# PROBABILITY THEORY

MATH 154

## Unit 14: Weak law of large numbers

**14.1.** A sequence of random variables  $X_k$  defines a **stochastic process**. There does not need to be anything stochastic about it.  $X_k$  could be the same for example. Or take  $Y_k(x) = \cos(2\pi(x + k\sqrt{2}))$  on  $([0, 1], \mathcal{A}, dx)$ , which shows correlations. The cumulative process  $S_n = X_1 + X_2 + \dots + X_n$  has a chance to converge if we scale it:  $S_n/n$ . For example, if  $X_i$  is the outcome of a dice, then  $S_n/n$  is the average of all  $n$  dice. Experience tells us that it is the average of the distribution:  $m = (1 + 2 + 3 + 4 + 5 + 6)/6 = 21/6 = 3.5$ . Under which conditions can we justify this? We will see later that even in the just mentioned cos process, we still have  $\frac{1}{n} \sum_{k=1}^n Y_k \rightarrow 0 = E[Y_k]$ .

**14.2.** The weak law of large numbers holds for pairwise uncorrelated random variables. This is a remarkably weak assumption.

**Theorem 1** (Golden Theorem). *Assume  $X_i \in \mathcal{L}^2$  are pairwise uncorrelated, have a common mean  $E[X_i] = m$  and  $M = \sup_n \frac{1}{n} \sum_{i=1}^n \text{Var}[X_i] < \infty$ . Then  $\frac{S_n}{n} \rightarrow m$  in probability.*

*Proof.* Since  $\text{Var}[X + Y] = \text{Var}[X] + \text{Var}[Y] + 2 \cdot \text{Cov}[X, Y]$  and  $X_n$  are pairwise uncorrelated, we get  $\text{Var}[X_n + X_m] = \text{Var}[X_n] + \text{Var}[X_m]$  and by induction  $\text{Var}[S_n] = \sum_{i=1}^n \text{Var}[X_i]$ . Using linearity, we obtain  $E[S_n/n] = m$  and

$$\text{Var}\left[\frac{S_n}{n}\right] = E\left[\frac{S_n^2}{n^2}\right] - \frac{E[S_n]^2}{n^2} = \frac{\text{Var}[S_n]}{n^2} = \frac{1}{n^2} \sum_{i=1}^n \text{Var}[X_i].$$

The right hand side converges to zero for  $n \rightarrow \infty$ . With Chebychev's inequality we obtain

$$P\left[\left|\frac{S_n}{n} - m\right| \geq \epsilon\right] \leq \frac{\text{Var}\left[\frac{S_n}{n}\right]}{\epsilon^2} \leq \frac{M}{n\epsilon^2}.$$

□

**14.3.** As an application in analysis, this leads to a constructive proof of a **theorem of Weierstrass** which states that polynomials are dense in the space  $C[0, 1]$  of all continuous functions on the interval  $[0, 1]$ .

**Theorem 2.** *For every  $f \in C[0, 1]$ , the **Bernstein polynomials***

$$B_n(x) = \sum_{k=1}^n f\left(\frac{k}{n}\right) \binom{n}{k} x^k (1-x)^{n-k}$$

*converge uniformly to  $f$ . If  $f(x) \geq 0$ , then also  $B_n(x) \geq 0$ .*

*Proof.* For  $x \in [0, 1]$ , let  $X_n$  be a sequence of independent  $\{0, 1\}$ -valued random variables with mean value  $x$ . In other words, we take the probability space  $(\{0, 1\}^{\mathbb{N}}, \mathcal{A}, \mathbb{P})$  defined by  $\mathbb{P}[\omega_n = 1] = x$ . Since  $\mathbb{P}[S_n = k] = \binom{n}{k} x^k (1-x)^{n-k}$ , we can write  $B_n(x) = \mathbb{E}[f(\frac{S_n}{n})]$ . We estimate with  $\|f\| = \max_{0 \leq x \leq 1} |f(x)|$

$$\begin{aligned} |B_n(x) - f(x)| &= |\mathbb{E}[f(\frac{S_n}{n})] - f(x)| \leq \mathbb{E}[|f(\frac{S_n}{n}) - f(x)|] \\ &\leq 2\|f\| \cdot \mathbb{P}[|\frac{S_n}{n} - x| \geq \delta] \\ &\quad + \sup_{|x-y| \leq \delta} |f(x) - f(y)| \cdot \mathbb{P}[|\frac{S_n}{n} - x| < \delta] \\ &\leq 2\|f\| \cdot \mathbb{P}[|\frac{S_n}{n} - x| \geq \delta] + \sup_{|x-y| \leq \delta} |f(x) - f(y)|. \end{aligned}$$

The second term in the last line is called the **continuity module** of  $f$ . It converges to zero for  $\delta \rightarrow 0$ . By the Chebychev inequality and the proof of the weak law of large numbers, the first term can be estimated from above by

$$2\|f\| \frac{\text{Var}[X_i]}{n\delta^2},$$

a bound which goes to zero for  $n \rightarrow \infty$  because the variance satisfies  $\text{Var}[X_i] = x(1-x) \leq 1/4$ .  $\square$

**14.4.** In the weak law of large numbers, we only assumed the random variables to be uncorrelated. Under the stronger condition of independence and the moment assumption  $X^4 \in \mathcal{L}^1$ , the convergence can be accelerated:

**Theorem 3** (Complete golden theorem). *Assume  $X_i \in \mathcal{L}^4$  have common expectation  $\mathbb{E}[X_i] = m$  and satisfy  $M = \sup_n \|X\|_4 < \infty$ . If  $X_i$  are independent, then  $S_n/n \rightarrow m$  in probability. We even have complete convergence:  $\sum_{n=1}^{\infty} \mathbb{P}[|\frac{S_n}{n} - m| \geq \epsilon]$  converges for all  $\epsilon > 0$ .*

*Proof.* We can assume without loss of generality that  $m = 0$ . Because the  $X_i$  are independent, we get

$$\mathbb{E}[S_n^4] = \sum_{i_1, i_2, i_3, i_4=1}^n \mathbb{E}[X_{i_1} X_{i_2} X_{i_3} X_{i_4}].$$

Again by independence, a summand  $\mathbb{E}[X_{i_1} X_{i_2} X_{i_3} X_{i_4}]$  is zero if an index  $i = i_k$  occurs alone, it is  $\mathbb{E}[X_i^4]$  if all indices are the same and  $\mathbb{E}[X_i^2] \mathbb{E}[X_j^2]$ , if there are two pairwise equal indices. Since by Jensen's inequality  $\mathbb{E}[X_i^2]^2 \leq \mathbb{E}[X_i^4] \leq M$  we get

$$\mathbb{E}[S_n^4] \leq nM + n(n-1)M.$$

Use now the Chebyshev-Markov inequality with  $h(x) = x^4$  to get

$$\mathbb{P}[|\frac{S_n}{n}| \geq \epsilon] \leq \frac{\mathbb{E}[(S_n/n)^4]}{\epsilon^4} \leq M \frac{n + n^2}{\epsilon^4 n^4} \leq 2M \frac{1}{\epsilon^4 n^2}.$$

$\square$